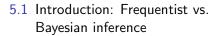


Matched line

GPS vertices

Rejected route (implied speed too fast)

5. Is the p value dead? Frequentist vs. Bayesian inference



5.2 General Methodics



- 5.4 Binary-Valued Quantities and Continuous Observations 5.4.1 Example: Map Matching
- 5.5 Continuous Quantities and Observations 5.5.1 Example: Gausian Priors and Observations
- 5.6 Conclusion

Econometrics Master's Course: Methods

▶ The classic **frequentist's** approach calculates the probability that the test function *T* is further away from *H*₀, (in the extreme range *E*_{data}) than the data realisation provided *H*₀ is marginally true:

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- ► The **Bayesian inference** tries to caculate what is actually interesting: The probability of *H*₀ given the data.
- ▶ If the unconditional or a-priori probabilities were known, this is easy using Bayes' theorem (abbreviating $T \in E_{data}$ as E_{data})

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- Example: *H*₀: "tomorrow is nice weather"
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P(B)

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Textbook case: binary variables \in { "true", "false" } (generalisations easy):

$$H_0: \beta = true, \quad \bar{H_0}: \beta = false, \quad B: \hat{\beta} = true; \quad \bar{B}: \hat{\beta} = false$$

$$P(H_0|B) = \frac{P(B|H_0)P(H_0)}{P(B)}$$

- H_0 : person is infected; B: person is tested positive
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 - Sensitivity $P(B|H_0) = 95\%$ $P(\bar{B}|H_0) = 5\%$
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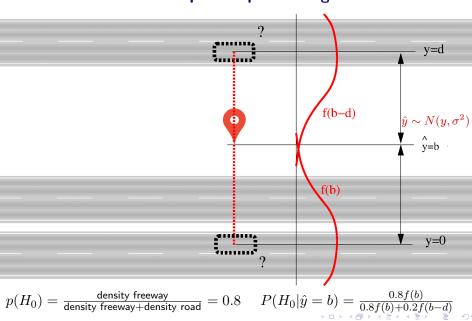
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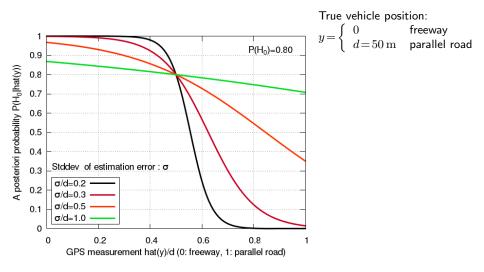
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Example: Map matching

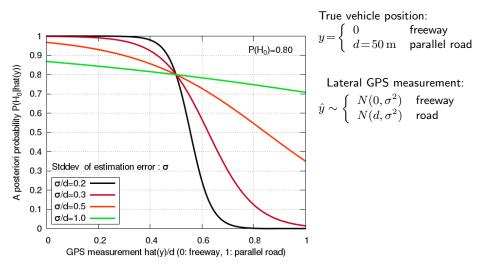


Map matching II

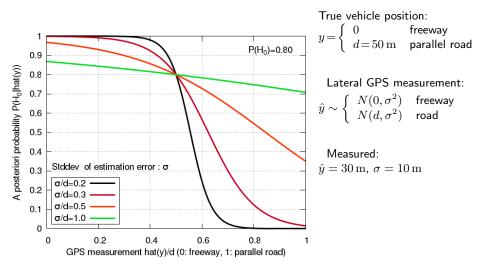


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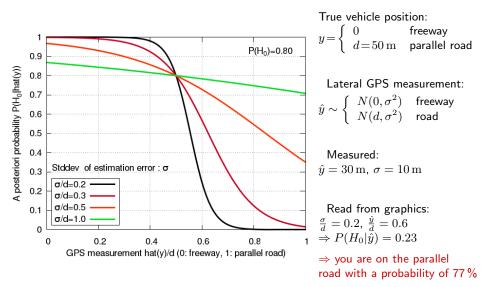
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Notice that the denominator is just the convolution $[f * h] \underset{(a) > b}{\text{at } \hat{\beta} = b}$

- -

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- ! We assume known variance, so $T=(\hat{\beta}-\beta_0)/\sigma_b\sim N(0,1).$ For $H_0\colon\beta\leq\beta_0$ we have

$$p = 1 - \Phi(t_{data})$$
$$= 1 - \Phi\left(\frac{b - \beta_0}{\sigma_b}\right)$$
$$\Phi\left(\frac{b - \beta_0}{\sigma_b}\right) = 1 - p$$
$$\frac{b - \beta_0}{\sigma_b} = \Phi^{-1}(1 - p)$$
$$b = \beta_0 + \sigma_b \Phi^{-1}(1 - p)$$

- ? Show that, on the previous slide, $eta_0=\sigma_B\Phi^{-1}(P(H_0))$
- ! We have $P(H_0) = P(\beta \le \beta_0) = \Phi\left(\frac{\beta_0}{\sigma_A}\right)$, so $\Phi^{-1}(P(H_0)) = \beta_0/\sigma_\beta$.

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$$\frac{b - \beta_0}{\sigma_b} = \Phi^{-1}(1 - p)$$

$$b = \beta_0 + \sigma_b \Phi^{-1}(1 - p)$$

- ? Show that, on the previous slide, $\beta_0 = \sigma_\beta \Phi^{-1}(P(H_0))$
- ! We have $P(H_0) = P(\beta \le \beta_0) = \Phi\left(rac{\beta_0}{\sigma_A}\right)$, so $\Phi^{-1}(P(H_0)) = \beta_0/\sigma_{\beta}$.

- ? Show that, on the previous slide, $b = \beta_0 + \sigma_b \Phi^{-1}(1-p)$
- ! We assume known variance, so $T=(\hat{\beta}-\beta_0)/\sigma_b\sim N(0,1).$ For $H_0\colon\beta\leq\beta_0$ we have

$$p = 1 - \Phi(t_{data})$$

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$$\mu = b rac{\sigma_{eta}^2}{\sigma_{eta}^2 + \sigma_b^2}$$

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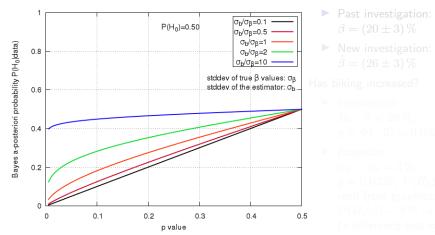
we have
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! Answer to the second question, $\sigma_{\beta} \ll \sigma_b$:

we have $\mu \to 0$, $\sigma \to \sigma_{\beta}$, $P(H_0|\hat{\beta}) = \Phi(\beta/\sigma_{\beta}) = P(H_0)$ ~

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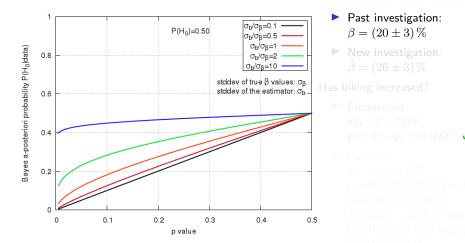
Bayesian inference for a Gaussian prior distribution 1: $P(H_0)=0.5$



Example: Bike modal split eta

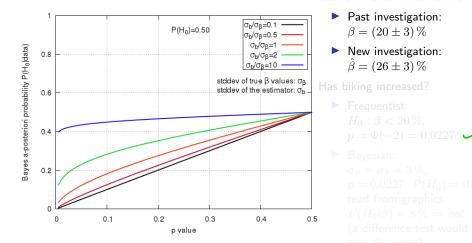
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Econometrics Master's Course: Methods



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Bayesian inference for a Gaussian prior distribution 1: $P(H_0)=0.5$



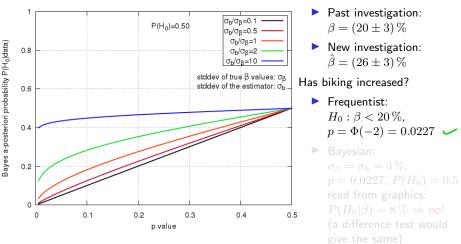
Bayesian inference for a Gaussian prior distribution 1: $P(H_0) = 0.5$

Past investigation: $\sigma_b/\sigma_\beta=0.1$ P(H₀)=0.50 $\beta = (20 \pm 3) \%$ $\sigma_{\rm b}/\sigma_{\rm B}=0.5$ $\sigma_b/\sigma_B=1$ Bayes a-posteriori probability P(H₀|data) New investigation: 0.8 $\sigma_b/\sigma_B=2$ $\hat{\beta} = (26 \pm 3) \%$ $\sigma_{\rm b}/\sigma_{\rm B}=10$ stddev of true β values: σ_β Has biking increased? stddev of the estimator: σ_b 0.6 Frequentist: 0.4 $p = \Phi(-2) = 0.0227$ 0.2 0 0 0.1 0.2 0.3 0.4 0.5 p value

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Bayesian inference for a Gaussian prior distribution 1: $P(H_0) = 0.5$

Example: Bike modal split β



Bayesian inference for a Gaussian prior distribution 1: $P(H_0)=0.5$

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p value

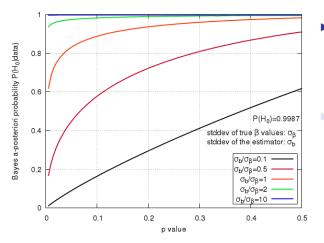
(a difference test would give the same)

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Example: Bike modal split β



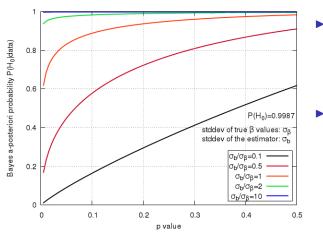
Bayesian inference for a Gaussian prior distribution 2: $P(H_0) = 0.9987$



• $\sigma_b \ll \sigma_\beta$ $\Rightarrow P(H_0|\hat{\beta}) \approx p$ \Rightarrow precise a-posteri information changes much.

• $\sigma_b \gg \sigma_\beta$ $\Rightarrow P(H_0|\hat{\beta}) \approx P(H_0)$ \Rightarrow fuzzy a-posteri data essentially give no information \Rightarrow a-priori probability nearly unchanged.

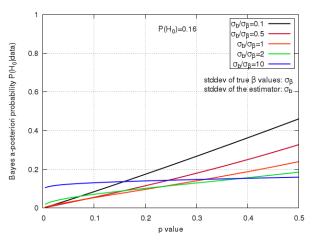
Bayesian inference for a Gaussian prior distribution 2: $P(H_0) = 0.9987$



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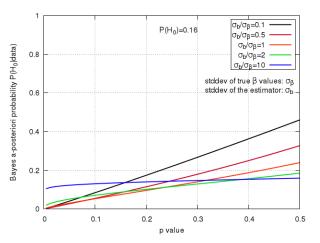
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Bayesian inference for a Gaussian prior distribution 3: $P(H_0) = 0.16$



Again, new data with $\sigma_b \ll \sigma_\beta$ gives much a-posteriori information (at least if p is significantly different from $P(H_0)$),

Bayesian inference for a Gaussian prior distribution 3: $P(H_0) = 0.16$



Again, new data with $\sigma_b \ll \sigma_\beta$ gives much a-posteriori information (at least if p is significantly different from $P(H_0)$),

New data with $\sigma_b \gg \sigma_\beta$ are tantamount to essentially no new information.

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► For discrete variables and measurements, we have the simple Bayes's calculations from elementary statistics → probability tree

Discrete variables and continuous measurements:

- If the measuring uncertainty is larger than the distance between possible discrete true values, then the a-priori probability matters
- If the uncertainty is much smaller, then the closest distance to the measurement matters
- The p value is completely mislading, even for bimodal continuous variables (vehicle not exactly in the middle of the right lane)
- Continuous variables and measurements:
 - The p value only gives a good estimate for the posterior probability $P(H_0|B)$ if (i) the prior distribution is unimodal, (ii) the measuring uncertainty is much smaller than the prior standard deviation, (iii) we have an interval null hypothesis
 - If the measuring uncertainty is much larger than the prior spread, the measurement hardly changes $P(H_0)$

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