5. Is the $p$ value dead? Frequentist vs. Bayesian inference
5.1 Introduction: Frequentist vs. Bayesian inference
5.2 General Methodics
5.3 Discrete Quantities and Observations 5.3.1 Example: Covid-19 Tests
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5.5.1 Example: Gausian Priors and Observations
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- The classic frequentist's approach calculates the probability that the test function $T$ is further away from $H_{0}$, (in the extreme range $E_{\text {data }}$ ) than the data realisation provided $H_{0}$ is marginally true:

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(iv) continuous sought-after quantity $\beta$ and continuous observation $\hat{\beta}$ (e.g., regression models)


### 5.3 Bayesian Inference for Discrete Quantities and Observations

Textbook case: binary variables $\in\{$ "true", "false" $\}$ (generalisations easy):

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H_{0}: \beta=\text { true }, \quad \bar{H}_{0}: \beta=\text { false }, \quad B: \hat{\beta}=\text { true } ; \quad \bar{B}: \hat{\beta}=\text { false }
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- $H_{0}$ after test negative: $P\left(H_{0} \mid \bar{B}\right)=P\left(\bar{B} \mid H_{0}\right) P\left(H_{0}\right) / P(\bar{B})=0.27 \%$


### 5.4 Bayesian Inference for Discrete Quantities and Continuous Observations

- Discrete quantity/parameter $\beta$ with the prior distribution $P\left(\beta=\beta_{j}\right)=p_{j}, \quad \sum_{j} p_{j}=1$
- Continuous measurement $\beta$ with a given distribution of density


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## - Assume $H_{0}$

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- Assume $H_{0}: \beta=\beta_{j_{0}}$ with $\beta_{j_{0}} \in\left\{\beta_{j}\right\}$ and the observation $B$ : $\hat{\beta} \in[b-\delta / 2, b+\delta / 2]$ with arbitrarily small $\delta$ :


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\Rightarrow \quad P\left(H_{0} \mid \hat{\beta}=b\right)=\frac{P\left(H_{0}\right) P\left(B \mid H_{0}\right)}{P(B)}=\frac{p_{j_{0}} f\left(b-\beta_{j_{0}}\right)}{\sum_{j} p_{j} f\left(b-\beta_{j}\right)}
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## Example: Map matching



$$
p\left(H_{0}\right)=\frac{\text { density freeway }}{\text { density freeway+density road }}=0.8 \quad P\left(H_{0} \mid \hat{y}=b\right)=\frac{0.8 f(b)}{0.8 f(b)+0.2 f(b-d)}
$$

## Map matching II



True vehicle position:
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Measured:
$\hat{y}=30 \mathrm{~m}, \sigma=10 \mathrm{~m}$

## Map matching II



# 5.5 Bayesian Inference for Continuous Quantities and Measurements 

- The quantity $\beta$ has the a-priori distribution density $h(\beta)$
- Unlike discrete quantities/parameters, $H_{0}$ needs to be an interval instead of a point (why?) $\Rightarrow P\left(H_{0}\right)$ and $P\left(B \mid H_{0}\right)$ are integrals over the values of $H_{0}$


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P\left(H_{0} \mid E_{\text {data }}\right)
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P\left(H_{0}\right) \rightarrow \int_{\beta \in H_{0}} h(\beta) \mathrm{d}(\beta) & \frac{\int_{\beta \in H_{0}} P\left(E_{\text {data }} \mid \beta\right) h(\beta) \mathrm{d} \beta}{\int_{\beta \in \mathbb{R}} P\left(E_{\text {data }} \mid \beta\right) h(\beta) \mathrm{d} \beta}
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\end{array}
$$

$P\left(E_{\text {data }} \mid \beta\right)$ is related to the $p$-value $P\left(E_{\text {data }} \mid \beta_{0} \in H_{0}^{*}\right)$ and also to the power function $\pi_{\alpha}(\beta)=P\left(R_{\alpha} \mid \beta\right) \quad\left[R_{\alpha}=\right.$ rejection region at $\left.\alpha\right]$

## Inference for a given measurement

Probability for $H_{0}$ based on a given realisation (measurement)
$\hat{\beta} \in B=[b-\delta / 2, b+\delta / 2]$ with arbitrarily small $\delta$ :

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\Rightarrow \quad P\left(H_{0} \mid B\right) & =\frac{\int_{\beta \in H_{0}} f(b-\beta) h(\beta) \mathrm{d} \beta}{\int_{\beta \in \mathbb{R}} f(b-\beta) h(\beta) \mathrm{d} \beta}
\end{aligned}
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Notice that the denominator is just the convolution $[f * h]$ at $\hat{\beta}=b$

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P\left(H_{0} \mid \hat{\beta}\right)=\Phi\left(\frac{\beta_{0}-\mu}{\sigma}\right), \quad \mu=b \frac{\sigma_{\beta}^{2}}{\sigma_{\beta}^{2}+\sigma_{b}^{2}}, \quad \sigma=\frac{\sigma_{\beta} \sigma_{b}}{\sqrt{\sigma_{\beta}^{2}+\sigma_{b}^{2}}}
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- When expressing the observation in terms of the $p$ value, $b=\beta_{0}+\sigma_{b} \Phi^{-1}(1-p)$ and $\beta_{0}$ in terms of $P\left(H_{0}\right)$, intervall null hypothesis for a single parameter $\beta$, any a-priori expectation $E(\beta)$, and any $H_{0}$ boundary value $\beta_{0}$


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- If $\sigma_{b}^{2} \ll \sigma_{\beta}^{2}$ and $H_{0}$ is an interval, we have $P\left(H_{0} \mid \hat{\beta}\right) \rightarrow p$ $\Rightarrow$ "ressurrection" of the $p$-value!


## Questions

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! We have $P\left(H_{0}\right)=P\left(\beta \leq \beta_{0}\right)=\Phi\left(\frac{\beta_{0}}{\sigma_{\beta}}\right)$, so $\Phi^{-1}\left(P\left(H_{0}\right)\right)=\beta_{0} / \sigma_{\beta}$.

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\sigma & =\sigma_{\beta} \sigma_{b} / \sqrt{\sigma_{\beta}^{2}+\sigma_{b}^{2}}
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&=
\end{aligned}\left(\frac{-\sigma_{b} \Phi^{-1}(1-p)}{\sigma_{b}}\right),
$$

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& =\Phi\left(-\Phi^{-1}(1-p)\right) \stackrel{\text { symm }}{=} \Phi\left(+\Phi^{-1}(p)\right)
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P\left(H_{0} \mid \hat{\beta}\right) & \rightarrow \Phi\left(\frac{\beta_{0}-b}{\sigma_{b}}\right) \quad \beta_{0}-b \stackrel{\text { in terms of } p}{=} \Phi\left(\frac{-\sigma_{b} \Phi^{-1}(1-p)}{\sigma_{b}}\right) \\
& =\Phi\left(-\Phi^{-1}(1-p)\right) \stackrel{\text { symm }}{=} \Phi\left(+\Phi^{-1}(p)\right) \stackrel{\text { def quantile }}{=} \underset{=}{=}
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## Questions II

? Show that, if the variance of the prior distribution is much larger than that of the measurement, we have $P\left(H_{0} \mid \hat{\beta}\right) \rightarrow p$ and, if it is much smaller, we have $P\left(H_{0} \mid \hat{\beta}\right) \rightarrow P\left(H_{0}\right)$
! Answer to the first question, $\sigma_{\beta} \gg \sigma_{b}$ :
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! Answer to the second question, $\sigma_{\beta} \ll \sigma_{b}$ :
we have $\mu \rightarrow 0, \sigma \rightarrow \sigma_{\beta}, P\left(H_{0} \mid \hat{\beta}\right)=\Phi\left(\beta / \sigma_{\beta}\right)=P\left(H_{0}\right)$

## Bayesian inference for a Gaussian prior distribution 1: $P\left(H_{0}\right)=0.5$



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- Past investigation:
$\beta=(20 \pm 3) \%$


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give the same)


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Has biking increased?
- Frequentist:

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\begin{aligned}
& H_{0}: \beta<20 \% \\
& p=\Phi(-2)=0.0227
\end{aligned}
$$

- Bayesian:

give the same)


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Example: Bike modal split $\beta$

- Past investigation:

$$
\beta=(20 \pm 3) \%
$$

- New investigation:
$\hat{\beta}=(26 \pm 3) \%$
Has biking increased?
- Frequentist:

$$
\begin{aligned}
& H_{0}: \beta<20 \% \\
& p=\Phi(-2)=0.0227
\end{aligned}
$$

- Bayesian:

$$
\sigma_{\beta}=\sigma_{b}=3 \%
$$

$$
p=0.0227, P\left(H_{0}\right)=0.5
$$ read from graphics:

$P\left(H_{0} \mid \hat{\beta}\right)=8 \% \Rightarrow$ no!
(a difference test would give the same)

## Bayesian inference for a Gaussian prior distribution 2: $P\left(H_{0}\right)=0.9987$


$>\sigma_{b} \ll \sigma_{\beta}$
$\Rightarrow P\left(H_{0} \mid \hat{\beta}\right) \approx p$
$\Rightarrow$ precise a-posteri
information changes much.
$\sigma_{b}$
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## Bayesian inference for a Gaussian prior distribution 3: $P\left(H_{0}\right)=0.16$



Again, new data with $\sigma_{b} \ll \sigma_{\beta}$ gives much a-posteriori information (at least if $p$ is significantly different from $P\left(H_{0}\right)$ ),

## Bayesian inference for a Gaussian prior distribution 3: $P\left(H_{0}\right)=0.16$



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New data with $\sigma_{b} \gg$ $\sigma_{\beta}$ are tantamount to essentially no new information.

### 5.6 Conclusion

- For discrete variables and measurements, we have the simple Bayes's calculations from elementary statistics $\rightarrow$ probability tree
- Discrete variables and continuous measurements:
- If the measuring uncertainty is larger than the distance between possible discrete true values, then the a-priori probability matters


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The $p$ value is completely mislading, even for bimodal continuous variables (vehicle not exactly in the middle of the right lane)


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- Continuous variables and measurements:
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- If the measuring uncertainty is much larger than the prior spread, the measurement hardly changes $P\left(H_{0}\right)$

